## Topics in Algorithmic Game Theory and Economics

#### Pieter Kleer

Max Planck Institute for Informatics (D1) Saarland Informatics Campus

February 3, 2020

**Lecture 11 Prophet Inequalities**

# Online selection

Consider

- Finite set of elements  $E = \{e_1, \ldots, e_m\}.$
- Weight function  $w : E \to \mathbb{R}_{\geq 0}$ .
- $\textsf{Collection of feasible subsets } \mathcal{F} \subseteq 2^\mathsf{E} = \{\mathcal{S} : \mathcal{S} \subseteq \mathsf{E}\}.$

Elements arrive one by one in unknown order  $\sigma = (\sigma(1), \ldots, \sigma(m))$ .

### Online selection problem with initial  $S = \emptyset$

For  $i = 1, \ldots, m$ , upon arrival of element  $\sigma(i)$ :

- Weight  $\mathsf{w}_{\sigma(i)}$  is revealed.
- **•** Decide (irrevocably) whether to select or reject  $\sigma(i)$ , where selecting is only allowed if  $S + \sigma(i) \in \mathcal{F}$ .

*Goal: Select subset S*  $\in$  *F maximizing w*(*S*) =  $\sum_{e \in S} w(e)$ *.* 

## **Bayesian setting**

*With adversarial arrival order*

# Bayesian setting

*Instead of making assumption on arrival order (uniform random), we make assumption on the (unknown) weights of the elements.*

In Bayesian setting, we have for every element *i* a (non-negative) probability distribution  $X_i$  :  $\mathbb{R}_{\geq 0} \to [0, 1].$ 

- $\bullet$  Distributions  $X_i$  are independent from each other.
- Weight  $w_i$  of element  $e_i$  is sample from  $X_i$ .

**Online procedure for set system**  $\mathcal{F} = (E, \mathcal{I})$ : Set  $S = \emptyset$ .

For every *i*, a realization  $w_i \sim X_i$  is generated.

• All realizations  $w_i$  are shown to the adversary.

• For 
$$
i = 1, \ldots, m
$$
:

- Adversary chooses  $\sigma(i) \in E,$  and reveals it and its weight  $w_i.$
- Online algorithm A decides whether to accept or reject  $\sigma(i)$ , where acceptance is only allowed if  $S + \sigma(i) \in \mathcal{F}$ .

# Probability distributions

Very roughly speaking, there are two main types of probability distributions: *continuous* and *discrete*.

A non-negative discrete random variable *X* is given by function  $g : \mathbb{N} \to [0,1]$  with

$$
\sum_{i=0}^{\infty} g(i) = 1.
$$

### Example

Suppose we have a fair die with six sides. Then  $g(i) = \frac{1}{6}$  for  $i = 1, \ldots, 6$  and  $g(i) = 0$  otherwise.

A non-negative continuous random variable *X* is given by density function  $f : \mathbb{R}_{\geq 0} \to \mathbb{R}_{\geq 0}$  with  $\int^{\infty}$ 0  $f(x)dx=1$ . It then holds that  $\mathbb{P}(X \leq z) = \int^{z}$ 0 *f*(*x*)*dx*.

$$
\mathbb{P}(X \leq z) = \int_0^z f(x) dx
$$

The function *f*(*x*) models how the probability mass is spread out.



#### Example

Consider the uniform distribution over the interval [a, b] with  $0 \le a \le b$ . Then  $f(x) = \frac{1}{b-a}$ .

### Remark

All the results we discuss today hold for both continuous and discrete distributions, but sometimes need slightly different arguments.

**Online procedure for set system**  $\mathcal{F} = (E, \mathcal{I})$ : Set  $S = \emptyset$ .

- For every *i*, a realization  $w_i \sim X_i$  is generated.
	- All realizations *w<sup>i</sup>* are shown to the adversary.
- For  $i = 1, ..., m$ :
	- Adversary chooses  $\sigma(i) \in E,$  and reveals it and its weight  $w_i.$
	- Online algorithm A decides whether to accept or reject  $\sigma(i)$ , where acceptance is only allowed if  $S + \sigma(i) \in \mathcal{F}$ .
- Algorithm A may use (in step *i*) information revealed so far, as well as the distributions *X<sup>i</sup>* of all elements.

### About the adversary

In general, we assume to have an all-knowing, adaptive adversary

- Can choose which element to present in step *i*, based on
	- Choices of online algorithm in steps 1, . . . , *i* − 1.
	- Realizations of all elements (including those that have not arrived).

Adversary is non-adaptive if order is fixed after seeing all realizations.

#### Example

Let  $E = \{e_1, e_2\}$  of which we may select at most one element. Let 1  $> \epsilon, \delta > 0$ , and assume that  $\frac{1}{\epsilon} > 1 + \delta$ . Distributions are given by:

$$
w_1 \sim X_1 = \begin{cases} \frac{1}{\epsilon} & \text{with probability } \epsilon \\ 0 & \text{with probability } 1 - \epsilon \end{cases}
$$
(1)  

$$
w_2 \sim X_2 = \begin{cases} 1 + \delta & \text{with probability } 1. \end{cases}
$$
(2)

Note that  $\mathbb{E}[X_1] = \frac{1}{\epsilon} \times \epsilon + 0 \times (1 - \epsilon) = 1$  and  $\mathbb{E}[X_2] = 1 + \delta$ .

**•** If arrival order would be  $(e_1, e_2)$ , simply observe realization  $w_1$ .

- If  $w_1 = 1/\epsilon$ , then select  $e_1$  (as  $\frac{1}{\epsilon} > 1 + \delta$ ).
- If  $w_1 = 0$ , reject  $e_1$  and select  $e_2$ .
- $\bullet$  Worst-case arrival order is  $(e_2, e_1)$ .
	- We don't know realization  $w_1$ , when deciding on element  $e_2$ .
	- Nevertheless, it is (intuitively) optimal to select  $e_2$ .
	- Why? Deterministic value  $w_2 = 1 + \delta > \mathbb{E}[X_1]$ .

 $\bullet$  In expectation (of  $X_1$ ), we cannot do better if we reject  $e_2$ .

• Performance objective is formalized next.

# Performance of online algorithm

Performance is measured against that of the prophet.

- Prophet gets to see all realizations *w<sup>i</sup>* ∼ *X<sup>i</sup>* after they are sampled.
- Computes (offline) subset *S* <sup>∗</sup> with max. weight

 $\mathsf{OPT}(w_1,\ldots,w_m):=w(\mathcal{S}^*)=\textit{max}_{\mathcal{S}\in\mathcal{F}}\sum_{e\in\mathcal{S}}w_e.$ 

• Expected weight for prophet is

$$
\text{OPT} = \mathbb{E}_{(y_1,\ldots,y_m)\sim X_1\times\cdots\times X_m}\left[\text{OPT}(y_1,\ldots,y_m)\right].
$$

Expected weight of (deterministic) algorithm  $A$  is

$$
\text{ALG} = \mathbb{E}_{(\textbf{y}_1,\ldots,\textbf{y}_m) \sim \textbf{X}_1 \times \cdots \times \textbf{X}_m} \left[ \min_{\sigma} \textbf{w}(\mathcal{A}(\sigma, \textbf{y}_1,\ldots,\textbf{y}_m)) \right].
$$

- With  $w(A(\sigma, y_1, \ldots, y_m))$  weight of set outputted by A.
- We assume to have a worst-case arrival order here.

For 
$$
0 < \alpha < 1
$$
, algorithm A is  $\alpha$ -approximation if

 $ALG > \alpha \cdot OPT$ .

This is called a prophet inequality.

### Example (cont'd)

 $E = \{e_1, e_2\}$  with following distributions. Let  $1 > \epsilon, \delta > 0$ , and assume that  $\frac{1}{\epsilon} > 1 + \delta$ . Let

$$
w_1 \sim X_1 = \begin{cases} \frac{1}{\epsilon} & \text{with probability } \epsilon \\ 0 & \text{with probability } 1 - \epsilon \end{cases}
$$
 (3)

$$
w_2 \sim X_2 = \{ 1 + \delta \text{ with probability 1.} \tag{4}
$$

#### **What can prophet get?**

$$
\mathsf{OPT}(w_1,w_2)=\max\{w_1,w_2\}=\left\{\begin{array}{ll} \frac{1}{\epsilon}&\text{with probability }\epsilon\\ 1+\delta&\text{with probability }1-\epsilon\end{array}\right..
$$

Then  $\mathbb{E}_{\mathsf{OPT}(\mathsf{y}_1, \mathsf{y}_2)}[\mathsf{max}_i|\mathsf{y}_i] = \frac{1}{\epsilon} \times \epsilon + (1+\delta) \times (1-\epsilon) \to 2$  as  $\epsilon, \delta \to 0$ .

#### Optimal algorithm  $\mathcal A$  is to select  $e_2$  (again, think about it).

- $\textsf{Worst-case order is } (\textit{\textbf{e}}_2, \textit{\textbf{e}}_1) \text{ with } \mathbb{E}_{(\textit{\textbf{y}}_1, \textit{\textbf{y}}_2)}[\textit{\textbf{w}}(\mathcal{A}(\sigma, \textit{\textbf{y}}_1, \textit{\textbf{y}}_2))] = 1.$
- I.e., optimal algorithm only half as bad as prophet ( $\alpha = \frac{1}{2}$  $\frac{1}{2}$ ).
- Also shows that, in general, we cannot hope for  ${\cal A}$  with  $\alpha>\frac{1}{2},$ already in setting where we can select at most one (out of two)  $\blacksquare$ elements.  $\blacksquare$

# **Selecting single element**

Prophet Inequality with  $\alpha = \frac{1}{2}$ 2

# Selecting single item

Krengel, Sucheston and Garling (1978) show there is a prophet inequality with  $\alpha = \frac{1}{2}$  $\frac{1}{2}$ .

- Simple algorithm was given by Samuel-Cahn (1984):
	- Set threshold  $T$  to be the median of distribution  $X_{\sf max} = \max_i X_i.$
	- Select first element *e<sup>i</sup>* whose realized *w<sup>i</sup>* ∼ *X<sup>i</sup>* exceeds threshold.

Median of distribution *X* is value *m* such that

$$
\mathbb{P}(X < m) \leq \frac{1}{2} \quad \text{and} \quad \mathbb{P}(X > m) \leq \frac{1}{2}.
$$

For continuous distributions, the median is the "middle value" of the distribution.

Example

Suppose we have uniform distribution over (continuous) interval [*a*, *b*]. Then  $m = \frac{a+b}{2}$  $\frac{+D}{2}$ .

As an alternative to Samuel-Cahn's median-based threshold, Kleinberg and Weinberg (2012) gave another threshold-based algorithm.

*Extends to case where multiple elements may be selected under matroid constraint.*

### KW-algorithm for (unknown) arrival order  $\sigma$

Let  $X_i$  be the distribution from which element  $e_i$ 's weight is drawn.

• Set threshold

$$
T = \frac{\mathbb{E}[\max_j X_j]}{2}.
$$

For  $i=1,\ldots,m$ : If  $w_{\sigma(i)}\geq 7$ , select  $\sigma(i)$  and STOP.

### Theorem (Kleinberg and Weinberg, 2012)

*The KW-algorithm selects an element e*<sup>∗</sup> *with the property that*

$$
\mathbb{E}_{X_1,\ldots,X_m}[w(e^*)]\geq \frac{1}{2}\cdot \mathbb{E}[\max_j X_j].
$$

$$
\mathbb{E}_{X_1,...,X_m}[w(e^*)] \geq \frac{\mathbb{E}[\max_j X_j]}{2} \ \ (=T)
$$

Proof: Let  $\tau \in \{1, \ldots, m\}$  be (random) step in which element is select, and let  $X<sub>\tau</sub>$  be the (random) weight of the selected element, i.e., it holds that

$$
\mathbb{E}[X_{\tau}]=\mathbb{E}_{X_1,...,X_m}[w(e^*)]
$$

• Assume w.l.o.g. that  $\sigma = (e_1, \ldots, e_m)$ . It holds that

$$
\mathbb{E}[X_{\tau}]=\int_0^T \mathbb{P}[X_{\tau}>x]dx+\int_T^{\infty}\mathbb{P}[X_{\tau}>x]dx
$$

when all distributions *X<sup>i</sup>* are continuous.

• See background material for discrete version of this claim:

$$
\mathbb{E}[X] = \sum_{k=0}^{\infty} \mathbb{P}[X \geq k].
$$

$$
\mathbb{E}[X_{\tau}]=\int_0^T \mathbb{P}[X_{\tau}>x]dx+\int_T^{\infty} \mathbb{P}[X_{\tau}>x]dx, \quad T=\frac{\mathbb{E}[\max_j X_j]}{2}
$$

Let  $p = P$ [max*<sub>i</sub>*  $X_i \geq T$ ].

 $\bullet$  1 – *p* is probability that we do not select anything.

• For any  $i = 1, \ldots, m$ , probability that we have not selected an element in step *i* is then at least  $1 - p$ .

It is not hard to see that

$$
\int_0^T \mathbb{P}[X_{\tau} > x] dx \geq \int_0^T \mathbb{P}[X_{\tau} > T] dx \geq \int_0^T p \cdot dx = pT.
$$
 (5)

Furthermore, for  $x > T$  it holds that

$$
\mathbb{P}[X_{\tau} > x] = \sum_{j=1}^{m} \mathbb{P}[X_{\tau} > x | \tau = j] \mathbb{P}[\tau = j]
$$
  
\n
$$
\geq (1 - p) \sum_{j=1}^{m} \mathbb{P}[X_{j} > x]
$$
  
\n
$$
\geq (1 - p) \mathbb{P}[\max_{j} X_{j} > x]
$$
 (union bound)

$$
\mathbb{E}[X_{\tau}] \geq pT + (1-p)\int_{T}^{\infty} \mathbb{P}[\max_{j} X_{j} > x]dx, \quad T = \frac{\mathbb{E}[\max_{j} X_{j}]}{2}
$$

Note that

$$
\mathbb{E}[\max_{j} X_{j}] = \int_{0}^{T} \mathbb{P}[\max_{j} X_{j} > x]dx + \int_{T}^{\infty} \mathbb{P}[\max_{j} X_{j} > x]dx = 2T
$$

by definition of  $\mathcal{T}.$  Since  $\int_0^{\mathcal{T}} \mathbb{P}[\mathsf{max}_j \, X_j > x] dx \leq \mathcal{T},$  it holds that

$$
\int_T^{\infty} \mathbb{P}[\max_j X_j > x] dx \geq T.
$$

Plugging this into the main inequality above gives

$$
\mathbb{E}[X_{\tau}] \geq pT + (1-p)T = T = \frac{\mathbb{E}[\max_j X_j]}{2}.
$$

This completes the proof.

### Theorem (Kleinberg and Weinberg, 2012)

*The KW-algorithm selects an element e*<sup>∗</sup> *with the property that*

$$
\mathbb{E}_{X_1,\ldots,X_m}[w(e^*)]\geq \frac{1}{2}\cdot \mathbb{E}[\max_j X_j].
$$

- Algorithm is optimal trade-off between weight of selected elements and probability of selecting an element.
	- Higher threshold would give better weight of selected element, but prob. that we can select one gets smaller.
	- Lower threshold would increase prob. of selecting element, but weight will be lower.
- Yields strategy proof online mechanism (in appropriate model).
	- Give item to first bidder exceeding threshold, and charge price *T*.
	- Similar to what we saw for secretary problem.

# **Matroid prophet inequality**

# Matroid prophet inequality

Selecting indep. set from matroid  $\mathcal{M} = (E, \mathcal{I})$  with arrival order  $\sigma$ .

Set  $S = \emptyset$ . For  $i = 1, \ldots, m$ , a realization  $w_i \sim X_i$  is generated. • All realizations  $w_i$  are shown to the adversary. • For  $i = 1, ..., m$ : Adversary chooses  $\sigma(i) \in E,$  and reveals it and its weight  $w_i.$ • Online algorithm A decides whether to accept or reject  $\sigma(i)$ , where acceptance is only allowed if  $S + \sigma(i) \in \mathcal{I}$ .

### Theorem (Kleinberg-Weinberg, 2012)

*There is an online algorithm* A *for selecting multiple elements subject to a matroid constraint (under adversarial arrival order), with*

> $\mathsf{ALG}(\mathcal{A})\geq \frac{1}{2}$  $\frac{1}{2}$   $\cdot$  OPT,

*where OPT* <sup>=</sup> <sup>E</sup>(*y*1,...,*ym*)∼*X*1×···×*X<sup>m</sup>* [*OPT*(*y*1, . . . , *ym*)] *is offline optimum.*

 $19/1$ 

## KW-algorithm for matroid constraint

Algorithm sets threshold in step *i* based on marginal contribution of  $\sigma(i)$ .

- Let  $y' = (y'_1, \ldots, y'_m) \ge 0$  be given weights, and let  $B'$  be a max. weight base under y'.
- For given independent set *S* ∈ I, we can augment *S* with  $\mathsf{elements}\;R(\mathcal{S})\subseteq\mathcal{B}'$  so that  $\mathcal{S}\cup R(\mathcal{S})$  is base of  $\mathcal{M}.$ 
	- Choose  $R$  so that  $y'(R)$  is maximized (among all choices for  $R$ ).



Assume that  $\sigma = (e_1, \ldots, e_m)$ .

### KW-algorithm with initial  $S = \emptyset$

For  $i = 1, \ldots, m$ : If  $S \cup \{e_i\} \in \mathcal{I}$  do the following.

**•** Set threshold

*T*<sub>*i*</sub> =  $\mathbb{E}_{y' \sim X_1 \times \cdots \times X_m}[y'(R(S)) - y'(R(S \cup \{e_i\}))].$ 

Set  $S \leftarrow S \cup \{e_i\}$  if  $w_i \geq T_i$ .

Roughly speaking, *T<sup>i</sup>* is expected gain of adding *e<sup>i</sup>* to *S*.

If revealed realization *w<sup>i</sup>* exceeds expected gain, add it to *S*.

In order to determine *T<sup>i</sup>* , we take expectation over all elements (and not just those that have not yet arrived).

**•** *T<sub>i</sub>* does not use realized weights *w*<sub>1</sub>, . . . , *w*<sub>*i*−1</sub> revealed so far.

*Computational remark: If the X<sup>i</sup> are discrete (with finite support), T<sup>i</sup> can be computed exactly (in possibly exponential time). For continuous distributions, usually approximation is needed (by means of repeatedly sampling vectors y' from*  $\times$ <sub>*i</sub>X<sub>i</sub> and computing average).*</sub>

## **Remarks**

### Theorem (Kleinberg and Weinberg, 2012)

KW-algorithm for matroids gives prophet inequality with  $\alpha = \frac{1}{2}$  $\frac{1}{2}$ .

- Result also extends to intersection of p matroid constraints, where one then gets  $\alpha = 1/(4p-2)$ .
- Can be used to model, e.g., setting where edges of bipartite graph arrive online (with known distributions).

### **Strategyproof mechanism?**

- For single element setting, conversion of respective KW-algorithm into strategyproof mechanism is easy.
- This is not the case for the matroid setting.

#### **Adaptive vs. non-adaptive** threshold-based algorithms.

- KW-algorithm is adaptive in the sense that threshold  $\mathcal{T}_i$  in step *i* depends on arrival order σ and elements *S* selected so far.
	- Does not necessarily yield strategyproof (online) mechanism.

# Non-adaptive threshold-based algorithms

A non-adaptive threshold-based algorithm sets threshold *T*(*e*) for every  $e \in E$  before start of the algorithm (independent of *i*).

It then selects *every* element whose weight exceeds the threshold (and that preserves independence).

Gives rise to so-called order-oblivious posted price mechanisms.

See Chawla, Goldner, Karlin and Miller (2020) for a (recent) algorithm for graphic matroids.

Interestingly, there exist matroid constraints for which *no* non-adaptive threshold-based algorithm can exist.

- Feldman, Svensson and Zenklusen (2020) give such an example for so-called gammoids.
- They show that one can hope at best for a prophet inequality with

$$
\alpha = \Omega\left(\frac{\log\log(m)}{\log(m)}\right).
$$

# **Beyond matroids**

For general downward-closed set systems, lower bound from last week also applies to Bayesian setting (with adversarial arrivals).

### Theorem (Babaioff et al. (2007), Rubinstein (2016))

*There is no randomized algorithm that, for every downward-closed set system*  $\mathcal{F} = (E, \mathcal{I})$  *with m* elements having known weight distribution, *obtains a prophet inequality with* α *better than*

$$
\alpha = \Omega\left(\frac{\log\log(m)}{\log(m)}\right)
$$

# **Selecting single element**

*Sample-based threshold*

## What prior information is needed?

Remember that the KW-algorithm for selecting a single item uses the threshold

$$
\mathcal{T}=\frac{\mathbb{E}[\max_j X_j]}{2}.
$$

Computing threshold requires full knowledge of the distributions *X<sup>i</sup>* .

Can be non-trivial depending on what the distributions look like.

*Does there exist an algorithm using less information?*

Turns out that it suffices to have one sample  $x_i$  from every  $X_i$ .

### Theorem (Rubinstein, Wang and Weinberg, 2020)

*Suppose we have one sample*  $x_i$  *form every*  $X_i$ *, and let*  $T = \max_j x_j$ *.* 

Selecting first element with  $w_i \geq T$  gives prophet inequality with  $\alpha = \frac{1}{2}$ 2 *.*

- Same guarantee as KW-algorithm.
- Algorithms only using single sample from every  $X_i$  will be called single-sample algorithms.

# Single-sample algorithms for matroid constraints

Azar, Kleinberg and Weinberg (2014) give single sample algorithms leading to constant-factor prophet inequalities for various matroid constraints.

- The high-level idea is to give a reduction to the secretary problem.
- Samples are used to mimic "observation phase" (Phase I).
	- Slightly stronger, order-oblivious secretary algorithm is needed.
	- An example is the  $\frac{1}{4}$ -approximation we saw in Homework 3.

### Theorem (Azar, Kleinberg and Weinberg, 2014 (informal))

*Every order-oblivious* α*-approximation for the secretary problem (with uniform random arrivals) gives rise to a single-sample prophet inequality with factor*  $\alpha$  *(for worst-case arrival order).* 

*Reduction also works for graphic matroid algorithm from last week.*

### Corollary (AKW, 2014)

*There is a single-sample*  $\alpha = \frac{1}{8}$ 8 *graphic matroid prophet inequality.*

# From single-sample prophets to secretaries

An algorithm (for adversarial arrival order  $\sigma$ ) with samples  $x_i$  from  $X_i$ : **Preprocessing:**

Set  $k = \frac{m}{2}$  $\frac{m}{2}$ , and select uniformly at random *k* samples from  ${x_1, \ldots, x_m}$ . Call the set of *k* samples

$$
\{y_{j_1},\ldots,y_{j_k}\}.
$$

### **Online:**

- For  $i = 1, \ldots, m$ , upon the arrival of  $\sigma(i)$ :
	- $\bullet$  If  $\sigma(i) \in \{j_1, \ldots, j_k\}$ , do nothing.
	- Otherwise, select  $\sigma(i)$  if  $w_i \ge \max\{y_{j_1}, \ldots, y_{j_k}\}.$

### Theorem (AKW, 2014)

*The above algorithm gives a single-sample prophet inequality with*  $\alpha = \frac{1}{4}$ 4 *for selecting one element.*

Proof uses the fact that both (offline) sample *x<sup>i</sup>* and (online) realization *w<sup>i</sup>* come from the same distribution *X<sup>i</sup>* .

## **Prophet inequalities for I.I.D. distributions**

## When all distributions *X<sup>i</sup>* are the same

Better prophet inequalities (than  $\alpha = \frac{1}{2}$  $\frac{1}{2}$ ) are possible when all distributions *X<sup>i</sup>* are the same.

The *X<sup>i</sup>* are independent and identically distributed (I.I.D.).

### Theorem (Correa et al., 2017)

*In case all the X<sup>i</sup> are I.I.D. there exists a prophet inequality with*  $\alpha \approx 0.745$  and this is best possible.

The algorithm has access to the weights revealed so far, and the common distribution *X*. What is possible when *X* is unknown?

### Theorem (Correa et al., 2018)

*In case the online algorithm only has access to weights revealed so far (but not to common distribution X ), there is a prophet inequality with*  $\alpha = \frac{1}{2}$ *e and this is best possible.*

# **Secretary prophet inequalities**

In the prophet secretary model, the elements in  $\{e_1, \ldots, e_m\}$ 

- **•** arrive in uniform random order
- with weight  $w_i$  drawn from known distribution  $X_i$  for  $i = 1, \ldots, m$ .

In this case, it is possible to obtain better results.

These results apply to the general setting with possibly non-I.I.D. distributions

### Theorem (Ehsani et al., 2018 (informal))

*There is a secretary prophet inequality with*  $\alpha = 1 - \frac{1}{e} \approx 0.63$  for *selecting multiple elements under a matroid constraint.*

### Theorem (Correa, Saona and Ziliotto, 2019)

*There is a secretary prophet inequality with*  $\alpha = 1 - \frac{1}{e} + \frac{1}{27} \approx 0.669$  for *selecting a single element.*

### **Overview**

*Have seen various online selection problems and models.*

Elements with unknown weights, but assumption on arrival order.

- Secretary problem
	- 1 *e* -approximation.
- Online bipartite matching (nodes on one side arriving online).
	- 1 *e* -approximation.
- Matroid secretary problem.
	- Open whether there is constant-factor approximation,
	- or possibly  $\frac{1}{e}$ -approximation.
- Most algorithms can be turned into online strategyproof mechanisms for selling items to (unit-)demand bidders.

Known weight distributions of elements, but adversarial arrival order.

- Prophet inequality with  $\alpha = \frac{1}{2}$  $\frac{1}{2}$  for selecting single element.
- Prophet inequality with  $\alpha = \frac{1}{2}$  $\frac{1}{2}$  for matroid constraint.
- Also saw some other models (e.g., single-sample settings).